# Statistical Curse of the Second Half Rank 

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and some recent developments
a problem from real life which can lead to a pretty much involved
combinatorics : ranking expectations in sailing boats regattas
example : the Spi Ouest France at la Trinité sur Mer (Brittany, each Easter)
involve a "large" number of identical boats $n_{b} \sim 100$
running a "large" number of races $n_{r} \sim 10=2,3$ races per day during 4 days

in each race each boat gets a rank $1 \leq \operatorname{rank} \leq 100$
no equal rank (no ex-aequo)
how to determine the final rank of a boat (and thus the winner) :

1) for each boat add its ranks in each race $\rightarrow$ its score $n_{t}$
here $n_{b}=100$ and $n_{r}=10 \Rightarrow 10 \leq n_{t} \leq 1000$
$n_{t}=10 \rightarrow$ lowest score always $1^{\text {rst }}$
$n_{t}=1000 \rightarrow$ highest score always $100^{\text {th }}$
$n_{t}=10 \times 50=500 \rightarrow$ middle score
2 ) order the scores $\Rightarrow$ final rank :
the boat with lowest score $\Rightarrow$ winner $1^{\text {rst }}$
the next boat after the winner $\Rightarrow$ second $2^{\text {nd }}$
what is the problem?
for example consider the ranks of a given boat to be
$51,67,76,66,55,39,67,59,66,54 \rightarrow$ its score $n_{t}=600$
clearly this boat has a mean rank $\frac{600}{10}=60$
$\rightarrow$ on average it has been $60^{\text {th }}$
$\rightarrow$ one might naively expect its final rank to be around $60^{\text {th }}$ no way : its final rank will rather be around $70^{\text {th }} \rightarrow$ "curse" see Spi Ouest 2009 data :
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| 60 | 496.00 | $3 \mathrm{~J}$ <br> X. Bourrut Lacouture | 53.00 | 60.00 | 35.00 | 58.00 | 66.00 | (69.00) | 62.00 | 63.00 | 36.00 | 63.00 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 61 | 499.00 | Atout Nautisme M. Bolou | 52.00 | 1.00) | 67.00 | 42.00 | 48.00 | 26.00 | 91.00 | 49.00 | 91.00 | 33.00 |
| 62 | 500.00 | Bmw Sailing Cup $\mathrm{N}^{\circ} 8$ <br> B. Le Rossignol | 28.00 | 76.00 | 76.00 | 8.00) | 57.00 | 15.00 | 61.00 | 74.00 | 46.00 | 67.00 |
| 63 | 501.00 | Jeroboam Marine Lorient X. Bonvarlet | 55.00 | 64.00 | 52.00 | 65.00 | 49.00 | 53.00 | 52.00 | 60.00 | 51.00 | 73.00) |
| 64 | 509.00 | $\begin{aligned} & \text { Icam-Olac } \\ & \text { A. Dal } \end{aligned}$ | 58.00 | 68.00 | 45.00 | 70.00 | 64.00 | 48.00 | 60.00 | 61.00 | 5.00) | 35.00 |
| 65 | 509.00 | Bmw Sailing Cup N ${ }^{\circ} 11$ <br> R. Lebohec | 68.00 | 55.00 | 65.00 | 57.00 | (91.00) | 40.00 | 55.00 | 65.00 | 60.00 | 44.00 |
| 66 | 510.00 | Ste Morbihannaise de Navigation <br> H. Dubois | 46.00 | 62.00 | 38.00 | 76.00 | (91.00) | 37.00 | 73.00 | 48.00 | 72.00 | 58.00 |
| 67 | 515.00 | Bmw Sailing Cup N ${ }^{\circ} 12$ <br> M. Dolle | (91.00) | 57.00 | 68.00 | 61.00 | 29.00 | 61.00 | 54.00 | 58.00 | 67.00 | 60.00 |
| 68 | 518.00 | J' Marine - Marine Lorient <br> G. Lautredou | 22.00 | 65.00 | 56.00 | 67.00 | 54.00 | (75.00) | 66.00 | 57.00 | 63.00 | 68.00 |
| 69 | 521.00 | Cholet <br> C. Bore | 63.00 | 59.00 | 51.00 | (68.00) | 61.00 | 63.00 | 64.00 | 53.00 | 48.00 | 59.00 |
| 70 | 524.00 | J -Venture M. Le Borgne | 51.00 | 67.00 | (76.00) | 66.00 | 55.00 | 39.00 | 67.00 | 59.00 | 66.00 | 54.00 |
| 71 | 543.00 | Bmw Sailing Cup $\mathrm{N}^{\circ} 2$ <br> O. Tarle | 50.00 | 74.00 | 50.00 | 64.00 | 39.00 | 72.00 | 68.00 | 70.00 | 56.00 | 1.00) |
| 72 | 553.00 | Penac'h B. Jaud | 66.00 | 61.00 | (73.00) | 72.00 | 56.00 | 65.00 | 56.00 | 56.00 | 57.00 | 64.00 |
| 73 | 560.00 | Ymir Junior H. Schilling | 45.00 | 66.00 | 64.00 | 56.00 | 53.00 | 44.00 | 68.00 | 73.00 | 1.00) | 91.00 |
| 74 | 564.00 | Art \& Stamps G. Le Baud | 77.00 | 70.00 | (88.00) | 71.00 | 72.00 | 77.00 | 59.00 | 40.00 | 58.00 | 40.00 |
| 75 | 571.00 | Marine Lorient P. Coindreau | 47.00 | (91.00) | 49.00 | 77.00 | 52.00 | 67.00 | 75.00 | 77.00 | 62.00 | 65.00 |
| 76 | 608.00 | Denis Pelfresne N. Barre | 46.00 | 63.00 | 78.00 | 80.00 | 70.00 | (85.00) | 74.00 | 66.00 | 61.00 | 70.00 |
| 77 | 630.00 | J'Mini <br> A. Ponsar | (91.00) | 73.00 | 81.00 | 83.00 | 32.00 | 82.00 | 63.00 | 67.00 | 74.00 | 75.00 |
| 78 | 645.00 | Jade Hisse Cn Pornic 1 <br> R. Romano | 54.00 | 78.00 | 75.00 | 75.00 | 67.00 | 73.00 | 79.00 | (91.00) | 64.00 | 80.00 |
| 79 | 653.00 | Mazda <br> G. Tarin | 70.00 | 77.00 | 72.00 | 62.00 | 69.00 | 76.00 | (83.00) | 72.00 | 76.00 | 79.00 |
| 80 | 655.00 | Ldt | 57.00 | 56.00 | 83.00 | 69.0 | (91.00) | 78.00 | 70.00 | 91.00 | 73.00 | 8.0 |

NB :
$51,67,(76), 66,55,39,67,59,66,54$
implies that the highest rank (76) is not taken into account
$\Rightarrow 9$ races : 51, 67, ,66, 55,39,67,59,66,54 $\rightarrow$ score $n_{t}=524$
$\Rightarrow$ mean rank $\frac{524}{9}=58$
on average $58^{\text {th }} \rightarrow 70^{\text {th }}$ even worse
a qualitative explanation of this "curse" is simple:
given the ranks of the boat : 51,67,76,66,55,39,67,59,66,54
assume that the ranks of the other boats are random variables with uniform distribution
random ranks : a good assumption if the crews are more or less equally worthy (which is in part the case)
since no ex aequo it means :
ranks of the other boats $=$ a random permutation
in the first race : random permutation of $(1,2,3, \ldots, 50,52, \ldots, 100)$
in the second race : random permutation of $(1,2,3, \ldots, 66,68, \ldots, 100)$
each race is obviously independent from the others
$\rightarrow$ a score is a sum of 10 independent random variables

10 is already a large number in probability calculus :
$\rightarrow$ Central Limit Theorem applies
$\rightarrow$ scores are random variables with gaussian probability density centered around the middle score $10 \times 50=500$
gaussian distribution $\Rightarrow$ a lot a boats with scores packed around 500
if the score of a boat is $>500$
its final rank is pushed upward from its mean rank

$$
\Rightarrow \text { statistical "curse" }
$$

on the contrary if the score of a boat is $<500$ its final rank is pushed downward from its mean rank

$$
\Rightarrow \text { statistical "blessing" }
$$

write things more precisely : namely given the score $n_{t}$ of a boat what is the probability distribution $P_{n_{t}}(m)$ for its final rank to be $m$ ?
a complication : $P_{n_{t}}(m)$ does not depend only on the score $n_{t}$ of the boat but also on its ranks in each race
for example : $n_{r}=3, n_{b}=3$ with a boat with score $n_{t}=6$
it is very easy to check by complete enumeration that $P_{6=2+2+2}(m) \neq P_{6=1+2+3}(m)$ (distributions are similar but different)
$\rightarrow$ a simplification : consider $n_{b}$ boats with random ranks
i.e ranks $=$ random permutation of $\left(1,2,3, \ldots, n_{b}\right)$
$\oplus$ an additional/virtual boat specified only by its score $n_{t}$
$\rightarrow$ same question : given the score $n_{t}$ of a virtual boat what is the probability distribution $P_{n_{t}}(m)$ for its final rank to be $m$ ?
$\rightarrow$ almost the same but simpler
call $n_{i, k}$ rank of the boat $i$ in a given race $k\left(1 \leq i \leq n_{b}\right.$ and $\left.1 \leq k \leq n_{r}\right)$

$$
\left\langle n_{i, k}\right\rangle=\frac{1+n_{b}}{2}
$$

no ex-aequo in race $k: \Rightarrow$ the $n_{i, k}$ 's are correlated random variables

$$
\begin{gathered}
\text { sum rule } \quad \sum_{i=1}^{n_{b}} n_{i, k}=1+2+3+\ldots+n_{b}=\frac{n_{b}\left(1+n_{b}\right)}{2} \\
\left\langle n_{i, k} n_{j, k}\right\rangle-\left\langle n_{i, k}\right\rangle\left\langle n_{j, k}\right\rangle=\frac{1+n_{b}}{12}\left(n_{b} \delta_{i, j}-1\right)
\end{gathered}
$$

$$
n_{i, k} \Rightarrow \text { score of boat } i=\sum_{k=1}^{n_{r}} n_{i, k} \equiv n_{i} \text { and middle score }=n_{r} \frac{1+n_{b}}{2}
$$

large $n_{r}$ limit $\rightarrow$ Central Limit Theorem for correlated random variables $\Rightarrow$ joint density probability distribution

$$
\begin{gathered}
f\left(n_{1}, \ldots, n_{n_{b}}\right)= \\
\sqrt{2 \pi \lambda n_{b}}\left(\sqrt{\frac{1}{2 \pi \lambda}}\right)^{n_{b}} \delta\left(\sum_{i=1}^{n_{b}}\left(n_{i}-n_{r} \frac{1+n_{b}}{2}\right)\right) \exp \left[-\frac{1}{2 \lambda} \sum_{i=1}^{n_{b}}\left(n_{i}-n_{r} \frac{1+n_{b}}{2}\right)^{2}\right] \\
\lambda=n_{r} \frac{n_{b}\left(1+n_{b}\right)}{12}
\end{gathered}
$$

for a virtual boat with score $n_{t}$ :
$P_{n_{t}}(m)$ is the probability for $m-1$ boats among the $n_{b}$ 's to have a score $n_{i}<n_{t}$ and for the other $n_{b}-m+1$ 's to have a score $n_{i} \geq n_{t}$

$$
P_{n_{t}}(m)=\binom{n_{b}}{m-1} \int_{-\infty}^{n_{t}} d n_{1} \ldots d n_{m-1} \int_{n_{t}}^{\infty} d n_{m} \ldots d n_{n_{b}} f\left(n_{1}, \ldots, n_{n_{b}}\right)
$$

take also large number of boats limit $\rightarrow$ saddle point approximation to finally get $\langle m\rangle=$ cumulative probability distribution of a normal variable

$$
\begin{gathered}
\langle m\rangle=\frac{n_{b}}{\sqrt{2 \pi \lambda}} \int_{-\infty}^{\bar{n}_{t}} \exp \left[-\frac{n^{2}}{2 \lambda}\right] d n \\
\bar{n}_{t}=n_{t}-n_{r} \frac{\left(1+n_{b}\right)}{2} \\
n_{r} \leq n_{t} \leq n_{r} n_{b} \rightarrow-n_{r} \frac{n_{b}}{2} \leq \bar{n}_{t} \leq n_{r} \frac{n_{b}}{2}
\end{gathered}
$$


variance

$$
\frac{(\Delta m)^{2}}{n_{b}}=\frac{1}{\sqrt{2 \pi \lambda}} \int_{-\infty}^{\bar{n}_{t}} \exp \left[-\frac{n^{2}}{2 \lambda}\right] d n \frac{1}{\sqrt{2 \pi \lambda}} \int_{-\infty}^{-\bar{n}_{t}} \exp \left[-\frac{n^{2}}{2 \lambda}\right] d n-\frac{1}{2 \pi} \exp \left[-\frac{\bar{n}_{t}^{2}}{\lambda}\right]
$$


$\Delta \mathrm{m}$

now consider small number of races $n_{r}=2,3, \ldots$ and boats $n_{b}=1,2, \ldots$
$\Rightarrow$ combinatorics problem
the simplest case $n_{r}=2$ : like a "2-body" problem
$\Rightarrow$ exact solution for $P_{n_{t}}(m)$
how to proceed :
i) represent possible configurations of ranks in the two races by points on a $n_{b} \times n_{b}$ lattice

2 races $\leftrightarrow$ square lattice, 3 races $\leftrightarrow$ cubic lattice, ...
no ex aequo $\Rightarrow 1$ point per line and per column
in general for $n_{b}$ boats and $n_{r}$ races $\rightarrow\left(n_{b}!\right)^{n_{r}-1}$ such configurations
ii) enumerate the configurations with $m-1$ points below the diagonal $n_{t}$
$\Rightarrow$ final rank $m$

$$
\begin{array}{r}
n_{b}=6 \\
n_{t}-1=5
\end{array}
$$

a $m=3$ configuration
4

2

combinatorics (not easy) :
for $2 \leq n_{t} \leq 1+n_{b}$
$\Rightarrow P_{n_{t}}(m)=\left(1+n_{b}\right) \sum_{k=0}^{m-1}(-1)^{k}\left(1+n_{b}-n_{t}+m-k\right)^{n_{t}-1} \frac{\left(n_{b}-n_{t}+m-k\right)!}{k!\left(1+n_{b}-k\right)!(m-k-1)!}$
for $2+n_{b} \leq n_{t} \leq 2 n_{b}+1$ by symmetry $P_{n_{b}+1-k}\left(n_{b}+2-m\right)=P_{n_{b}+2+k}(m)$
for the middle score $n_{t}=2 \frac{1+n_{b}}{2}=1+n_{b}$

$$
\Rightarrow P_{n_{t}=1+n_{b}}(m)=\left(1+n_{b}\right) \sum_{k=0}^{m-1}(-1)^{k} \frac{(m-k)^{n_{b}}}{k!\left(1+n_{b}-k\right)!}
$$

```
Table[p[5, nt, m], {nt, 2, 6}, {m, 1, 6}]
```

$$
\begin{aligned}
& \left\{\{1,0,0,0,0,0\},\left\{\frac{4}{5}, \frac{1}{5}, 0,0,0,0\right\},\left\{\frac{9}{20}, \frac{1}{2}, \frac{1}{20}, 0,0,0\right\},\right. \\
& \left.\left\{\frac{2}{15}, \frac{11}{20}, \frac{3}{10}, \frac{1}{60}, 0,0\right\},\left\{\frac{1}{120}, \frac{13}{60}, \frac{11}{20}, \frac{13}{60}, \frac{1}{120}, 0\right\}\right\}
\end{aligned}
$$

$$
\text { Table }[p[n b, n t=1+n b, m] n b!,\{n b, 1,7\},\{m, 1, n b+1\}]
$$

$$
\{\{1,0\},\{1,1,0\},\{1,4,1,0\},\{1,11,11,1,0\},\{1,26,66,26,1,0\},
$$

$$
\{1,57,302,302,57,1,0\},\{1,120,1191,2416,1191,120,1,0\}\}
$$

$\rightarrow$ Eulerian numbers

$$
\begin{aligned}
& \alpha=\frac{I}{I(p-I)} \\
& b=\frac{p+I}{I .2(p-I)^{2}} \\
& g=\frac{p p+4 p+1}{1.2 .3(p-I)^{3}} . \\
& \delta=\frac{p^{3}+11 p^{2}+11 p+1}{1 \cdot 2 \cdot 3 \cdot 4(p-1)^{4}} \\
& \varepsilon=\frac{p^{4}+26 p^{3}+66 p^{2}+26 p+1}{1.2 \cdot 3 \cdot 4 \cdot 5(p-1)^{3}} \\
& \zeta=\frac{p^{5}+57 p^{4}+302 p^{3}+302 p^{2}+57 p+I}{x \cdot 2 \cdot 3 \cdot 4 \cdot 5 \cdot 6(p-x)^{6}} \\
& \eta=\frac{p^{6}+120 p^{3}+1191 p^{4}+2416 p^{3}+1191 p^{2}+120 p+1}{1.2 .3 \cdot 4 \cdot 5 \cdot 6.7(p-1)^{7}} \\
& \text { sec. } \\
& \text { Eulerian Polynomials } \\
& \frac{A_{n}(p) / p}{n!(p-1)^{n}} \quad(1 \leq n \leq 7)
\end{aligned}
$$

## Eulerian number $=$

the number of permutations of the numbers 1 to $n$ in which exactly $m$ elements are greater than the previous element (permutations with $m$ "ascents")

| $\boldsymbol{n}$ | $\boldsymbol{m}$ | $\quad$ Permutations |
| :--- | :--- | :--- |
| $\mathbf{1}$ | 0 | $(1)$ |
| 2 | 0 | $(2,1)$ |
|  | 1 | $(1,2)$ |
|  | 0 | $(3,2,1)$ |
| 3 | 1 | $(1,3,2)(2,1,3)(2,3,1)(3,1,2)$ |
| 2 | $(1,2,3)$ |  |
| $n$ | $=4$ | $(1,4,2,3) \rightarrow m=2$ |

generating function
$g[x, y]=\frac{e^{x}(-1+y)}{-e^{x y}+e^{x} y}$
Series [g[x, y], \{x, 0, 6\}]

$$
\begin{aligned}
& 1+x+\frac{1}{2}(1+y) x^{2}+\frac{1}{6}\left(1+4 y+y^{2}\right) x^{3}+\frac{1}{24}\left(1+11 y+11 y^{2}+y^{3}\right) x^{4}+ \\
& \frac{1}{120}\left(1+26 y+66 y^{2}+26 y^{3}+y^{4}\right) x^{5}+\frac{1}{720}\left(1+57 y+302 y^{2}+302 y^{3}+57 y^{4}+y^{5}\right) x^{6}+0[x]^{7}
\end{aligned}
$$

why Eulerian numbers should play a role here seems a mystery but : an other way to look at things by rewriting

$$
P_{n_{t}}(m)=\frac{1}{n_{b}!} \sum_{i=m}(-1)^{i+m} n_{n_{t}}(i)\left(1+n_{b}-i\right)!\binom{i-1}{m-1}
$$

$n_{n_{t}}(i)=$ Stirling partition numbers : count in how many ways can the numbers $\left(1,2, \ldots, n_{t}-1\right)$ be partitioned in $i$ groups
example $n_{t}=5: \rightarrow 1$ way to split the numbers $(1,2,3,4)$ into 1 group
$\rightarrow 7$ ways to split the numbers $(1,2,3,4)$ into 2 groups
$(1),(2,3,4) ;(2),(1,3,4) ;(3),(1,2,4) ;(4),(1,2,3) ;(1,2),(3,4) ;(1,3),(2,4) ;(1,4),(2,3)$
$\rightarrow 6$ ways to split the numbers $(1,2,3,4)$ into 3 groups
$(1),(2),(3,4) ;(1),(3),(2,4) ;(1),(4),(2,3) ;(2),(3),(1,4) ;(2),(4),(1,3) ;(3),(4),(1,2)$
$\rightarrow 1$ way to split the numbers $(1,2,3,4)$ into 4 groups

$$
n_{t}=5 \rightarrow 1,6,7,1
$$

why Stirling numbers should play a role here seems again a mystery they appear from graph counting considerations on the configuration lattice :
for example for $n_{t}=5$
consider all the points below the diagonal

$a \left\lvert\, \begin{aligned} & 1 \\ & 1\end{aligned}\right.$
b $\left\lvert\, \begin{aligned} & 1 \\ & 2\end{aligned}\right.$
C| $\begin{aligned} & 1 \\ & 3\end{aligned}$
d $\left.\right|_{1} ^{2}$
$e \left\lvert\, \begin{array}{ll}2 & f \\ 2 & \\ 1\end{array}\right.$


$$
\Rightarrow \quad 6, \quad 7,1
$$

so from graph counting
$n_{n_{t}+1}(i+1)=$ under the diagonal $n_{t}$ number of subgraphs with $i$ points fully connected
$\rightarrow$ recurrence relation :
either 0 point on the diagonal $n_{t}-1 \rightarrow n_{n_{t}}(i+1)\binom{n_{t}-1}{0}$
either 1 point on the diagonal $n_{t}-1 \rightarrow n_{n_{t}-1}(i)\binom{n_{t}-1}{1}$
either 2 points on the diagonal $n_{t}-1 \rightarrow n_{n_{t}-2}(i-1)\binom{n_{t}-1}{2}$
$\Rightarrow n_{n_{t}+1}(i+1)=\sum_{k=0}^{i} n_{n_{t}-k}(i+1-k)\binom{n_{t}-1}{k}$
$\Leftrightarrow$ recurrence relation for Stirling partition numbers
there is indeed a one to one correspondance between Stirling partition (in fact second class Stirling) numbers and Eulerian numbers

$$
\operatorname{Eulerian}[\mathrm{n}, \mathrm{k}]=\sum_{j=1}^{k+1}(-1)^{k-j+1}\binom{n-j}{n-k-1} j!\text { Stirling }[n, j]
$$

why all this?

$$
\begin{gathered}
2 \text { races : } P_{n_{t}}(m)=\frac{1}{n_{b}!} \sum_{i=m}(-1)^{i+m} n_{n_{t}}(i)\left(1+n_{b}-i\right)!\binom{i-1}{m-1} \\
\rightarrow n_{r} \text { races }: P_{n_{t}}(m)=\frac{1}{\left(n_{b}!\right)^{n_{r}-1}} \sum_{i=m}(-1)^{i+m} n_{n_{t}}\left(i, n_{r}\right)\left(1+n_{b}-i\right)!^{n_{r}-1}\binom{i-1}{m-1}
\end{gathered}
$$

how to calculate $n_{n_{t}}\left(i, n_{r}\right) ? \rightarrow$ work in progress

